

## Regular Article

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# Application of Finite Mixture to Characterise Degraded *Gmelina arborea* Roxb Plantation in Omo Forest Reserve, Nigeria

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## Abstract

The use of single component distribution to describe the irregular stand structure of degraded forest often lead to bias. Such biasness can be overcome by the application of finite mixture distribution. Therefore, in this study, finite mixture distribution was used to characterise the irregular stand structure of the *Gmelina arborea* plantation in Omo forest reserve. Thirty plots, ten each from the three stands established in 1984, 1990 and 2005 were used. The data were pooled per stand and fitted. Four finite mixture distributions including normal mixture, lognormal mixture, gamma mixture and Weibull mixture were considered. The method of maximum likelihood was used to fit the finite mixture distributions to the data. Model assessment was based on negative loglikelihood value ( $-2\ln L$ ), Akaike information criterion (AIC), Bayesian information criterion (BIC) and root mean square error (RMSE). The results showed that the mixture distributions provide accurate and precise characterisation of the irregular diameter distribution of the degraded *Gmelina arborea* stands. The  $-2\ln L$ , AIC, BIC and RMSE values ranged from -715.233 to -348.375, 703.926 to 1433.588, 718.598 to 1451.334 and 3.003 to 7.492, respectively. Their performances were relatively the same. This approach can be used to describe other irregular forest stand structures, especially the multi-species forest.

**Key Words:** finite mixture distribution, gamma mixture, normal mixture, lognormal mixture, Weibull mixture.

## Introduction

The diameters of trees, and mainly its structure characterise any forest stand very well. Diameter structure is a very significant stand characteristic based on which we can evaluate the growth and volume production of stand as well as the structure of assortments, maturity, etc. Diameter structure may be expressed by frequency of tree diameters that quantify their distribution in diameter classes (Petras et al. 2010). Several probability distribution functions have been used to describe the diameter and/or height distributions in natural and forest plantation (manmade) for-

ests e.g. beta, exponential, gamma, Johnson's SB, logit-logistic, lognormal, normal, Weibull distributions etc.

Modelling the diameter distributions of pure monospecific even-aged stands is a simple issue and this might be achieved by using any of the aforesaid statistical distributions. However, the diameter structure of forest is often affected by disturbance caused by natural disaster (e.g. wind throw) and/or anthropogenic activities (e.g. thinning, illegal logging, bush burning etc.) (Tsogt and Lin 2012). This may result to irregular, bimodal or heavily skewed diameter distributions. Under such condition, a single univariate distribution function may not uniquely de-

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scribe the diameter structure of the forest stand. The use of single component distribution may lead to oversimplification of the forest stand structure (Maltamo et al. 2000; Liu et al. 2002). To capture this irregularity, a finite mixture model has been recommended for modelling stand with two or more modes (bimodal or multimodal) diameter distributions (Bailey and Dell 1973; Zhang et al. 2001).

Finite mixture modelling is an approach applied in the classification of observations and to model unobserved heterogeneity. In this approach, the observed data are generally assumed to come from unobserved subpopulations called classes or groups; as such, mixtures of probability densities e.g. normal, Weibull etc. are used to model the variable of interest. One of the most reliable and preferred methods of fitting mixtures of distribution is the method of maximum likelihood through the expectation maximization (EM) algorithm. This algorithm has “reliable convergence properties, however, convergence can be slow if the component of the mixture are not well separated and the sample size is small” (Izenman and Sommer 1988).

The concept of finite mixture modelling was first used in tree diameter distributions by Zhang et al. (2001), Liu et al. (2002) and Zasada and Cieszewski (2005). Zhang et al. (2001) used a finite mixture of two Weibull functions to model the diameter distributions of uneven-aged stands. The authors reported a better result with mixture Weibull compared to a single Weibull and exponential distributions. Liu et al. (2002) also used a finite mixture model to characterise the diameter distributions of mixed-species forest stands. Zasada and Cieszewski (2005) showed the applicability of finite mixture distribution approach for characterising tree diameter distributions by natural social class in pure even-aged stand. Some of the recent studies on finite mixture modelling in tree diameter distribution include Zhang and Liu (2006), Podlaski (2010a, 2010b, 2017), Jaworski and Podlaski (2012), Tsogt and Lin (2012), Lui et al. (2014). In these studies, the authors found finite mixture modelling as a promising method for modelling irregular, bimodal or multimodal diameter distributions.

To date, most distribution studies have relied on the use of a single component distribution to model irregular stand structure of degraded forest. The *Gmelina arborea* stands in Omo forest reserve have suffered from severe illegal logging activities, bush burning and agricultural encroach-

ment (Ogana et al. 2017). This has exposed the forest estate to wind-throw. In consequence, result to irregular stand structure. Therefore, the main objective of this study is to apply the finite mixture distributions to characterise the irregular diameter distribution of the degraded *Gmelina arborea* stands in Omo forest reserve, Nigeria.

## Materials and Methods

### Data

The data used for this study came from the *Gmelina arborea* plantation in Omo Forest Reserve. Its lies between Latitudes 6°35' and 7°05'N and Longitudes 4°19' and 4°10'E of Ogun State, Nigeria and occupies an area of 130,500 ha (Ogana et al. 2017). A total of 30 sample plots of 0.04 ha from three distinct stands established in 1984, 1990 and 2005 were used. Ten plots were laid in each stand. The descriptive statistics of the inventory data are presented in Table 1.

### Mixture distribution

The finite mixture distribution for a random continuous variable x (e.g. tree diameter, height etc.) can be expressed as:

$$f(x,p) = \sum_{j=1}^k p_j f_j(x) = p_1 f_1(x) + p_2 f_2(x) + \dots + p_k f_k(x) \quad (1)$$

Where  $f_j(x)$  is the probability density function (pdf) of the  $j$ th individual component distribution,  $p$  is the mixing proportion of the individual components in the mixture distribution and must satisfy the condition:  $0 \leq p_j \leq 1$  and

**Table 1.** Descriptive statistics of the data set

| Stand | Tree variable | Statistics |      |     |                    |
|-------|---------------|------------|------|-----|--------------------|
|       |               | Mean       | Max  | Min | Standard deviation |
| 1984  | Dbh (cm)      | 18.5       | 42.9 | 5.8 | 9.14               |
| N=257 | Height (m)    | 13.7       | 42.7 | 1.6 | 6.59               |
| 1990  | Dbh (cm)      | 22.2       | 40.8 | 5.7 | 9.35               |
| N=240 | Height (m)    | 15.5       | 26.3 | 5.9 | 5.13               |
| 2005  | Dbh (cm)      | 15.3       | 34.1 | 5.9 | 6.26               |
| N=139 | Height (m)    | 8.8        | 14.9 | 3.1 | 2.43               |

$p_k = \sum_{j=1}^k p_j$ . In this study, the normal mixture, lognormal mixture, gamma mixture and Weibull mixture were used to characterise the irregular tree diameter distributions of the *Gmelina arborea* stands. The normal mixture consists of two component normal distributions. the lognormal mixture also has two component lognormal distributions. Similarly, the gamma mixture and Weibull mixture have two component gamma and two components Weibull distributions, respectively. The component distributions are expressed as:

$$\text{Normal pdf } f(x) = \frac{\exp(-\frac{1}{2}(\frac{x-\mu}{\sigma})^2)}{\sigma\sqrt{2\pi}} \quad (2)$$

$$\text{Lognormal pdf } f(x) = \frac{\exp(-\frac{1}{2}(\frac{\ln x - \mu}{\sigma})^2)}{x\sigma\sqrt{2\pi}} \quad (3)$$

Where  $\mu$  and  $\sigma$  are the location and scale parameters of the normal and lognormal function,  $x$ =diameter at breast height (Dbh)

$$\text{Gamma pdf } f(x) = \frac{x^{\alpha-1}}{\beta^\alpha\Gamma(\alpha)}\exp(-\frac{x}{\beta}) \quad (4)$$

2-parameter Weibull pdf (Weibull 1951):

$$f(x) = \frac{\alpha}{\beta} (\frac{x}{\beta})^{\alpha-1} \exp[-(\frac{x}{\beta})^\alpha] \quad (5)$$

Where  $\alpha$  and  $\beta$  are the shape and scale parameters, respectively.  $\Gamma(\alpha)$  is the gamma function with parameters  $\alpha$  where  $\alpha > 0$ ,  $\beta > 0$  and  $\Gamma(\alpha) = (\alpha - 1)!$

The method of maximum likelihood through the expectation maximization (EM) algorithm was used to fit the mixture distributions to the data. The R software (R Core Team 2017) was used in this study. Due to anthropogenic activities in the study site, few trees were encountered in the plots across the 3 stands; as such, the plot data were pooled per stand and fitted. Model assessment was based on negative loglikelihood value, Akaike information criterion (AIC), Bayesian information criterion (BIC) and root mean square error (RMSE). The smaller the values are, the better the model.

## Result and Discussion

Finite mixture distributions have been used to described the diameter distributions of the disturbed *Gmelina arborea* stands in Omo Forest Reserve. The estimated parameters of the finite mixture distributions are presented in Table 2. The mixing parameter  $p$  gives the contribution of each component to the overall diameter distribution of the forest

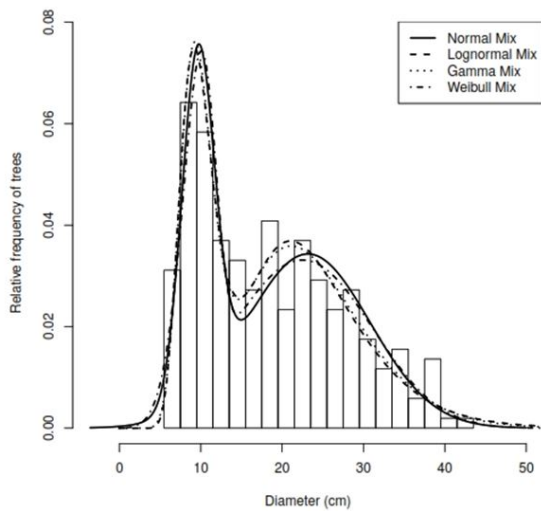
**Table 2.** Estimated parameters of the four finite mixture distributions for the three stands

| Stand | Distributions   | Parameters |            |            |       |            |            |
|-------|-----------------|------------|------------|------------|-------|------------|------------|
|       | 2 Normal mix    | $p_1$      | $\mu_1$    | $\sigma_1$ | $p_2$ | $\mu_2$    | $\sigma_2$ |
| 1984  |                 | 0.342      | 9.674      | 1.995      | 0.658 | 23.170     | 7.647      |
| 1990  |                 | 0.412      | 13.052     | 4.361      | 0.588 | 28.736     | 5.542      |
| 2005  |                 | 0.472      | 10.434     | 2.929      | 0.528 | 19.519     | 5.145      |
|       | 2 Lognormal mix | $p_1$      | $\mu_1$    | $\sigma_1$ | $p_2$ | $\mu_2$    | $\sigma_2$ |
| 1984  |                 | 0.390      | 2.265      | 0.223      | 0.610 | 3.139      | 0.300      |
| 1990  |                 | 0.526      | 2.647      | 0.421      | 0.474 | 3.385      | 0.164      |
| 2005  |                 | 0.393      | 2.235      | 0.286      | 0.607 | 2.896      | 0.274      |
|       | 2 Gamma mix     | $p_1$      | $\alpha_1$ | $\beta_1$  | $p_2$ | $\alpha_2$ | $\beta_2$  |
| 1984  |                 | 0.356      | 23.326     | 0.412      | 0.644 | 9.864      | 2.382      |
| 1990  |                 | 0.520      | 6.344      | 2.381      | 0.480 | 38.126     | 0.788      |
| 2005  |                 | 0.484      | 11.358     | 0.918      | 0.516 | 16.274     | 1.212      |
|       | 2 Weibull mix   | $p_1$      | $\alpha_1$ | $\beta_1$  | $p_2$ | $\alpha_2$ | $\beta_2$  |
| 1984  |                 | 0.315      | 5.555      | 10.311     | 0.685 | 3.154      | 25.385     |
| 1990  |                 | 0.399      | 3.190      | 14.586     | 0.601 | 5.584      | 30.747     |
| 2005  |                 | 0.415      | 3.983      | 11.479     | 0.585 | 3.634      | 20.677     |

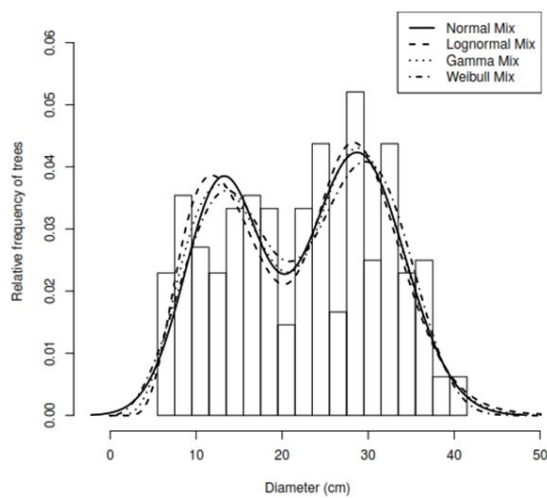
stand. This parameter adds up to 1 for each distribution for the individual stand. This is a major condition for finite mixture model. For example, the mixing parameter values of the 2 component normal mix for stand 1984 were 0.342 and 0.658 -which summed up to 1. This shows that the diameter data of the stands have two distinct groups. The number of components mixtures to include in a finite mixture distribution could be predetermined especially when other information on the data set are available. During for-

est inventory, trees are usually classified or grouped based on species, tree position etc. (Zasada and Cieszewski 2005). This information can be used to determine the number of component mixtures in the distributions. Jaworski and Podlaski (2012) noted that “the chance of lack of convergence during parameter estimation process increases with the number of components of the finite mixture model”. However, in this study, an iterative search was used to determine the optimum number of component mixture with the smallest fit indices. This method was also recommended by Zhang et al. (2001) and Liu et al. (2002).

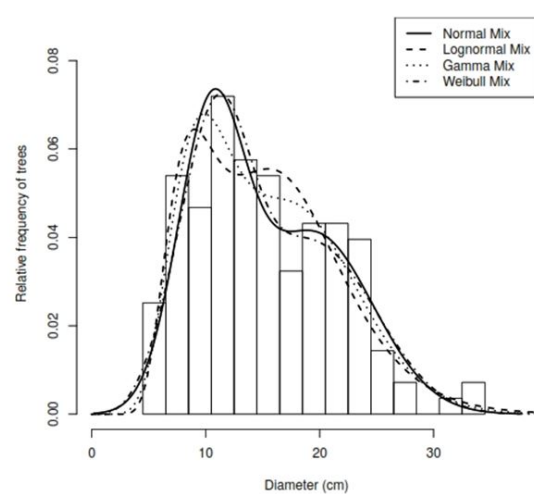
The diameter distributions of the three disturbed *Gmelina arborea* forest stands in Omo Forest Reserve are presented in Fig. 1-4. The graphs showed the observed relative frequency of trees with the fitted normal mix, lognormal mix, gamma mix and Weibull mix by diameter classes for stand 1984, 1990 and 2005 (Fig. 1-3). The mixture distributions described the irregular diameter structure of the forest stands very well compared to the single component distributions. For comparison, the single component of normal lognormal, gamma and Weibull distributions were fitted to the data (see Fig. 4-6). The graph showed that the single component distribution underestimated the frequency in the 10cm diameter class and overestimated in the 18 and 22 cm classes in stand 1984. In addition, the single component distributions resulted in the overall simplification of the forest stand which is in con-



**Fig. 1.** Observed relative frequency of trees per Dbh class and fitted mixture distributions for stand 1984.



**Fig. 2.** Observed relative frequency of trees per Dbh class and fitted mixture distributions for stand 1990.



**Fig. 3.** Observed relative frequency of trees per Dbh class and fitted mixture distributions for stand 2005.

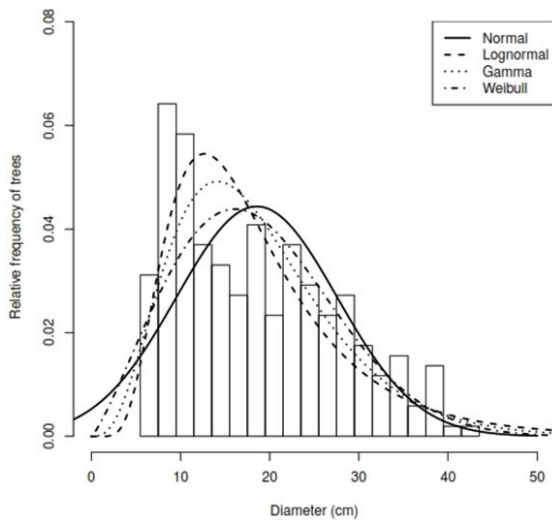


Fig. 4. Observed relative frequency of trees per Dbh class and fitted single component distributions for stand 1984.

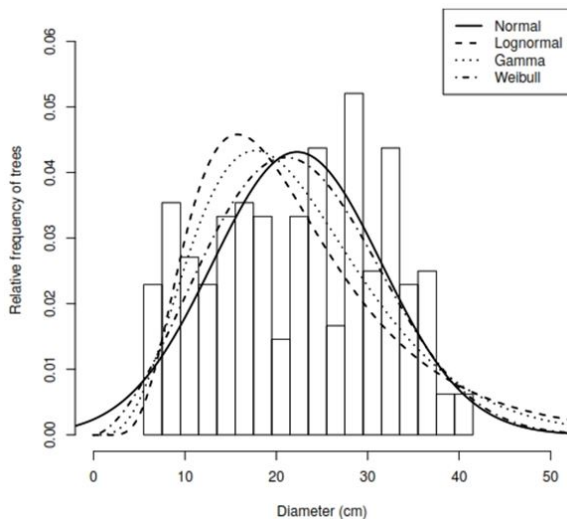


Fig. 5. Observed relative frequency of trees per Dbh class and fitted single component distributions for stand 1990.

trasts with reality. However, the mixture distributions fitted to the same stand revealed the irregularity in the forest stands. Similar observation was reported by Lui et al. (2002), Zhang and Liu (2006) and Lui et al. (2014). The irregularities in the structure of the *Gmelina arborea* stands can be attributed to the illegal logging activities and agricultural activities. This has been a major constraint of the forest sector, especially in Nigeria.

The negative loglikelihood ( $-\Lambda$ ), AIC, BIC and

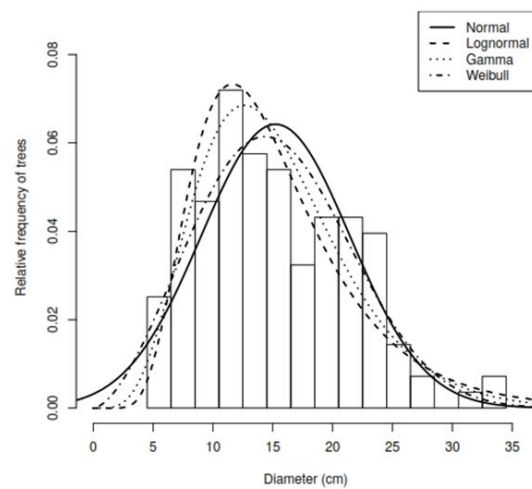


Fig. 6. Observed relative frequency of trees per Dbh class and fitted single component distributions for stand 2005.

RMSE values of the normal mix, lognormal mix, gamma mix and Weibull mix were relatively the same for the three stands (Table 3). The  $-\Lambda$ , AIC, BIC and RMSE values ranged from -715.233 to -348.375, 703.926 to 1433.588, 718.598 to 1451.334 and 3.003 to 7.492, respectively. The fit indices values for the single component distributions were larger than the mixture distributions (though not documented). This shows the inadequacy of the single component distributions in describing the irregular diameter distribution of the degraded forest stand. The study is in line with Liu et al. (2014) who reported lower RMSE value for the mixture Weibull distribution than the single component Weibull in their study on “modelling diameter distributions of mixed-species forest stands”. Similarly, Jaworski and Podlaski (2012) found a better performance with mixture gamma than the single component gamma distribution. It is obvious that finite mixture modelling remains the most effective and efficient technique that can be used to model multi-species, irregular or complex forest stands. The fitting method especially, maximum likelihood through the EM algorithm is relative simple, and can be achieved with different statistical tools.

In conclusion, the application of finite mixture of normal, lognormal, gamma and Weibull distributions provides more accurate and precise characterisation of the irregular diameter distribution of the degraded *Gmelina arborea* plantation; in consequence, the forest stand structure. This

**Table 3.** Fit indices for the mixture distribution

| Stand | Distribution  | $-\Lambda\Lambda$ | AIC      | BIC      | RMSE  |
|-------|---------------|-------------------|----------|----------|-------|
| 1984  | Normal mix    | 715.223           | 1440.447 | 1458.192 | 7.242 |
|       | Lognormal mix | 712.485           | 1434.969 | 1452.714 | 6.188 |
|       | Gamma mix     | 711.794           | 1433.588 | 1451.334 | 7.302 |
|       | Weibull mix   | 714.764           | 1439.529 | 1457.274 | 7.492 |
| 1990  | Normal mix    | 690.534           | 1391.067 | 1408.470 | 5.699 |
|       | Lognormal mix | 689.097           | 1388.194 | 1405.597 | 5.193 |
|       | Gamma mix     | 688.238           | 1386.477 | 1403.879 | 4.965 |
|       | Weibull mix   | 688.037           | 1386.073 | 1403.477 | 5.455 |
| 2005  | Normal mix    | 348.371           | 706.742  | 721.415  | 3.003 |
|       | Lognormal mix | 347.437           | 704.874  | 719.546  | 4.220 |
|       | Gamma mix     | 346.963           | 703.926  | 718.598  | 3.006 |
|       | Weibull mix   | 348.375           | 706.749  | 721.422  | 3.441 |

$-\Lambda\Lambda$  = negative loglikelihood (the larger the value in negative sign, the better the model).

approach can be used to describe other irregular forest stand structures, especially the multi-species forest in Nigeria.

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